## COMS6998-11: Homework 2 Solutions

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1. **Importance weighting and policy gradient.** For the first part, we have to combine to arguments we have seen previously: that "on policy" roll-outs with geometric stopping is unbiased, and that importance weighting is unbiased. The first part is shown by the following calculation (throughout we are conditioning on  $s_0 = s$ ,  $a_0 = a$ ):

$$\mathbb{E}\left[\left(\prod_{t=1}^{t_{\star}} \frac{\pi_2(a_t \mid s_t)}{\pi_1(a_t \mid s_t)}\right) \frac{r_{t^{\star}}}{1-\gamma}\right] = \sum_{T=0}^{\infty} \Pr[t^{\star} = T] \mathbb{E}\left[\left(\prod_{t=1}^{T} \frac{\pi_2(a_t \mid s_t)}{\pi_1(a_t \mid s_t)}\right) \frac{r_T}{1-\gamma}\right]$$
$$= \sum_{T=0}^{\infty} \mathbb{E}\left[\left(\prod_{t=1}^{T} \frac{\pi_2(a_t \mid s_t)}{\pi_1(a_t \mid s_t)}\right) \gamma^T r_T\right]$$

For the second part, let's consider just one of the terms above and expand the expectation over trajectories  $\tau_t = (s_0, a_0, s_1, a_1, \dots, s_T, a_T)$ . Here we let r(s, a) denote the expected reward from (s, a).

$$\mathbb{E}\left[\left(\prod_{t=1}^{T} \frac{\pi_{2}(a_{t} \mid s_{t})}{\pi_{1}(a_{t} \mid s_{t})}\right) \gamma^{T} r_{T}\right] = \sum_{\tau_{T}} \mathbb{P}^{\pi_{1}}[\tau_{t}] \left(\prod_{t=1}^{T} \frac{\pi_{2}(a_{t} \mid s_{t})}{\pi_{1}(a_{t} \mid s_{t})}\right) \gamma^{T} r(s_{T}, a_{T})$$

$$= \sum_{\tau_{T}} \left(\prod_{t=1}^{T} P(s_{t} \mid s_{t-1}, a_{t-1}) \pi_{1}(a_{t} \mid s_{t})\right) \left(\prod_{t=1}^{T} \frac{\pi_{2}(a_{t} \mid s_{t})}{\pi_{1}(a_{t} \mid s_{t})}\right) \gamma^{T} r(s_{T}, a_{T})$$

$$= \sum_{\tau_{T}} \left(\prod_{t=1}^{T} P(s_{t} \mid s_{t-1}, a_{t-1}) \pi_{2}(a_{t} \mid s_{t})\right) \gamma^{T} r(s_{T}, a_{T})$$

$$= \sum_{\tau_{T}} \mathbb{P}^{\pi_{2}}[\tau_{t}] \gamma^{T} r(s_{T}, a_{T}).$$

Putting this together with the previously display, we obtain the result.

For the second part, the calculation is somewhat straightforward:

$$\frac{\pi_2(a \mid s)}{\pi_1(a \mid s)} = \frac{\exp(c(s, a))}{\sum_{a'} \pi_1(a' \mid s) \exp(c(s, a'))} \le \frac{\exp(c_\star)}{\exp(-c_\star) \sum_{a'} \pi_1(a' \mid s)} \le \exp(2c^\star).$$

A similar calculation applies for the other direction.

Finally for the third part, we focus on finding a deterministic quantity  $t_{\text{max}}$  such that  $t_{\star} < t_{\text{max}}$  with high probability. If this holds (formally, conditioned on  $t_{\star} \le t_{\text{max}}$ ), we know that  $\hat{Q}^{\pi_2}(s, a) \le \frac{\exp(2t_{\text{max}}c^{\star})}{1-\gamma} =: Q_{\text{max}}$  with probability 1, so we will have proved the result.

By a direct calculation

$$\Pr[t^* \ge T] = \sum_{\tau=T}^{\infty} (1 - \gamma) \gamma^{\tau} = \gamma^T (1 - \gamma) \sum_{\tau=0}^{\infty} \gamma^{\tau} = \gamma^T$$

Therefore  $t_{\text{max}} \ge \log(1/\delta)/\log(1/\gamma)$  suffices.

2. Tabular RL with generative models. Consider some policy  $\pi$  and some (s, a) pair. (For notation only we consider  $\pi$  to be deterministic but this is not essential.) First note that since rewards are in [0, 1] we have that  $Q_{M_2}^{\pi} \in [0, \frac{1}{1-\gamma}]$ . Then

$$\begin{split} |Q_{M_1}^{\pi}(s,a) - Q_{M_2}^{\pi}(s,a)| &= |\mathbb{E}_{R_1(s,a)}[r] + \gamma \mathbb{E}_{s' \sim P_1(s,a)} Q_{M_1}^{\pi}(s',\pi(s')) - \mathbb{E}_{R_2(s,a)}[r] - \gamma \mathbb{E}_{s' \sim P_2(s,a)} Q_{M_2}^{\pi}(s',\pi(s'))| \\ &\leq |\mathbb{E}_{R_1(s,a)}[r] - \mathbb{E}_{R_2(s,a)}[r]| + \gamma |\mathbb{E}_{s' \sim P_1(s,a)}[Q_{M_1}^{\pi}(s',\pi(s'))] - \mathbb{E}_{s' \sim P_2(s,a)}[Q_{M_2}^{\pi}(s',\pi(s'))]| \\ &\leq \varepsilon + \gamma \left| \mathbb{E}_{P_1(s,a)}[Q_{M_1}^{\pi}(s',\pi(s')) - Q_{M_2}^{\pi}(s',\pi(s'))] \right| + \gamma \left| (\mathbb{E}_{P_1(s,a)} - \mathbb{E}_{P_2(s,a)}) Q_{M_2}^{\pi}(s',\pi(s')) \right| \\ &\leq \varepsilon + \frac{\gamma \varepsilon}{1 - \gamma} + \gamma \left| \mathbb{E}_{P_1(s,a)}[Q_{M_1}^{\pi}(s',\pi(s')) - Q_{M_2}^{\pi}(s',\pi(s'))] \right| \\ &\leq \varepsilon + \frac{\gamma \varepsilon}{1 - \gamma} + \gamma \max_{s,a} \left| Q_{M_1}^{\pi}(s,a) - Q_{M_2}^{\pi}(s,a) \right|. \end{split}$$

Here, the first equality uses the definitions of the Q functions. In the second we use the triangle inequality to separate the immediate reward from the next-step value functions. In the third line we use the assumed bound on the reward differences and we also add and subtract a "cross term" quantity:  $\mathbb{E}_{s'\sim P_1(s,a)}Q_{M_2}^{\pi}(s',\pi(s'))$ . This leads us to a one step error term as well as a recursive term. The one step term is bounded via  $|(\mathbb{E}_P - \mathbb{E}_Q)(f(x))| \leq \sup_x |f(x)| \cdot ||P - Q||_{\text{TV}}$ .

Now let  $\bar{s}, \bar{a}$  be the state-action pair that maximize the difference  $(\bar{s}, \bar{a}) = \operatorname{argmax}_{s,a} |Q_{M_1}^{\pi}(s, a) - Q_{M_2}^{\pi}(s, a)|$ . Then we have just showed that

$$|Q_{M_1}^\pi(\bar{s},\bar{a}) - Q_{M_2}^\pi(\bar{s},\bar{a})| \le \varepsilon + \frac{\gamma\varepsilon}{1-\gamma} + \gamma \left|Q_{M_1}^\pi(\bar{s},\bar{a}) - Q_{M_2}^\pi(\bar{s},\bar{a})\right|.$$

We can re-arrange this to obtain a bound for all state-action pairs.

$$|Q_{M_1}^{\pi}(s,a) - Q_{M_2}^{\pi}(s,a)| \le |Q_{M_1}^{\pi}(\bar{s},\bar{a}) - Q_{M_2}^{\pi}(\bar{s},\bar{a})| \le \frac{\varepsilon}{1-\gamma} + \frac{\gamma\varepsilon}{(1-\gamma)^2} \le \frac{2\varepsilon}{(1-\gamma)^2}.$$

For the second part, consider a single (s, a) pair and obtain n samples  $\{(r_i, s_i')\}_{i=1}^n$  from the sampling oracle. Then by Hoeffding's inequality the empirical reward  $\bar{R}(s, a) = \frac{1}{n} \sum_{i=1}^{n} r_i$  satisfies (w.p.  $1 - \delta$ )

$$\left| \bar{R}(s,a) - \mathbb{E}_{R(s,a)}[r] \right| \lesssim \sqrt{\frac{\log(1/\delta)}{n}}.$$

Meanwhile, by Bernstein's inequality the empirical transition probability  $\hat{P}(s' \mid s, a) = \frac{1}{n} \sum_{i=1}^{n} \mathbf{1}\{s'_i = s'\}$  satisfies

$$\left| \hat{P}(s' \mid s, a) - P(s' \mid s, a) \right| \lesssim \sqrt{\frac{P(s' \mid s, a) \log(1/\delta)}{n}} + \frac{\log(1/\delta)}{n}.$$

Therefore, taking a union bound over all choices of s' we have

$$\|\hat{P}(s,a) - P(s,a)\|_{\text{TV}} \le \frac{1}{2} \sum_{s'} \left| \hat{P}(s' \mid s,a) - P(s' \mid s,a) \right|$$
$$\lesssim \sum_{s'} \left( \sqrt{\frac{P(s' \mid s,a) \log(S/\delta)}{n}} + \frac{\log(S/\delta)}{n} \right)$$
$$\lesssim \sqrt{\frac{S \log(S/\delta)}{n}} + \frac{S \log(S/\delta)}{n}$$

With a union bound the above argument holds simultaneously for all SA pairs. Thus if we set  $n = O(S \log(SA/\delta)/\varepsilon^2)$  we have uniform approximation, using  $O(S^2A \log(SA/\delta)/\varepsilon^2)$  samples in total.

If we find the optimal policy in the approximate MDP  $\hat{M}$ , by an analysis similar to that for ERM, we have

$$J_M(\pi^*) \le J_{\hat{M}}(\pi^*) + \frac{2\varepsilon}{(1-\gamma)^2} \le J_{\hat{M}}(\hat{\pi}) + \frac{2\varepsilon}{(1-\gamma)^2}$$
$$\le J_M(\pi^*) + \frac{4\varepsilon}{(1-\gamma)^2}$$

Thus to obtain suboptimality  $\epsilon$  in total, we require

$$O\left(\frac{S^2 A \log(SA/\delta)}{(1-\gamma)^4 \epsilon^2}\right)$$

samples in total.

3. **Generative models for linear MDPs.** We use a recursive argument, similar to the proof of the simulation lemma:

$$\left| Q^{(T)}(s,a) - Q^{\star}(s,a) \right| \leq \left| \widehat{TV^{(T-1)}}(s,a) - TV^{(T-1)}(s,a) \right| + \left| TV^{(T-1)}(s,a) - Q^{\star}(s,a) \right|$$

Using the linear MDP property, there exists a  $\bar{w}$  such that  $\mathcal{T}V^{(T-1)}(s,a) = \langle \phi(s,a), \bar{w} \rangle$ , while we constructed the empirical backup to satisfy  $\widehat{\mathcal{T}V^{(T-1)}}(s,a) = \langle \phi(s,a), \hat{w} \rangle$ . Introducing the covariance matrix, we have

$$\begin{split} \left| \widehat{TV^{(T-1)}}(s,a) - \widehat{TV^{(T-1)}}(s,a) \right| &= \left| \left\langle \phi(s,a), \overline{w} - \hat{w} \right\rangle \right| \le \|\phi(s,a)\|_{\Sigma^{-1}} \cdot \|\overline{w} - \hat{w}\|_{\Sigma} \\ &\le \sqrt{d} \cdot \sqrt{\mathbb{E}_D \left[ (\widehat{TV^{(T-1)}}(s,a) - \widehat{TV^{(T-1)}}(s,a))^2 \right]} \\ &\le \sqrt{\frac{d\Delta \log(1/\delta)}{n}}. \end{split}$$

This takes care of the first term. The second term has a recursive form

$$\left| \mathcal{T}V^{(T-1)}(s,a) - Q^{\star}(s,a) \right| = \gamma \left| \mathbb{E}_{s' \sim P(s,a)} \left[ V^{(T-1)}(s') - V^{\star}(s') \right] \right|$$

Recall that  $V^{(T-1)}(s') = \max_{a'} Q^{(T-1)}(s',a')$  while  $V^{\star}(s') = \max_{a'} Q^{\star}(s',a')$ . We would like to obtain a difference in the two Q-functions on the same state-action pair as this will allow us to recurse the argument. For this we introduce the policy  $\tilde{\pi}(s) = \operatorname{argmax}_a \max\{Q^{(T-1)}(s,a), Q^{\star}(s,a)\}$ . This policy leads to the inequality

$$\left| V^{(T-1)}(s') - V^{\star}(s') \right| = \left| \max_{a'} Q^{(T-1)}(s', a') - \max_{a'} Q^{\star}(s', a') \right| \le \left| Q^{(T-1)}(s', \tilde{\pi}(s')) - Q^{\star}(s', \tilde{\pi}(s')) \right|$$

To see why this is true, suppose that  $\tilde{\pi}(s')$  is the greedy action w.r.t.,  $Q^{(T-1)}(s',\cdot)$ . Then

$$Q^{(T-1)}(s', \tilde{\pi}(s')) \ge Q^{\star}(s', \pi^{\star}(s')) \ge Q^{\star}(s', \tilde{\pi}(s')),$$

as desired. A similar argument holds in the other case. Therefore,

$$\gamma \left| \mathbb{E}_{s' \sim P(s,a)} \left[ V^{(T-1)}(s') - V^{\star}(s') \right] \right| \le \gamma \cdot \sup_{s,a} \left| Q^{(T-1)}(s,a) - Q^{\star}(s,a) \right|$$

Putting things together, we have

$$\sup_{s,a} |Q^{(T)}(s,a) - Q^{\star}(s,a)| \le \sqrt{\frac{d\Delta \log(1/\delta)}{n}} + \gamma \sup_{s,a} |Q^{(T-1)}(s,a) - Q^{\star}(s,a)|$$

We can apply the same argument on the last term on the right hand side and unrolling this gives

$$\sup_{s,a} |Q^{(T)}(s,a) - Q^{\star}(s,a)| \le \sum_{t=0}^{T} \gamma^t \sqrt{\frac{d\Delta \log(1/\delta)}{n}} + \frac{\gamma^T}{1-\gamma} \le \frac{1}{1-\gamma} \sqrt{\frac{d\Delta \log(1/\delta)}{n}} + \frac{\gamma^T}{1-\gamma}.$$

The final term uses the trivial bound that  $|Q^{(1)} - Q^*| \le \frac{1}{1-\gamma}$  simply because  $Q^*(s, a) \in [0, 1/(1-\gamma)]$ . For the second part, let the right hand side of the bound in part (a) be  $\varepsilon_Q$ . Then

$$\begin{split} J(\pi^{\star}) - J(\hat{\pi}) &= \mathbb{E}_{s_0} Q^{\star}(s_0, \pi^{\star}(s_0)) - Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) \\ &= \mathbb{E}_{s_0} Q^{\star}(s_0, \pi^{\star}(s_0)) - Q^{\star}(s_0, \hat{\pi}(s_0)) + Q^{\star}(s_0, \hat{\pi}(s_0)) - Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) \\ &\leq \mathbb{E}_{s_0} Q^{\star}(s_0, \pi^{\star}(s_0)) - Q^{(T)}(s_0, \pi^{\star}(s_0)) + Q^{(T)}(s_0, \hat{\pi}(s_0)) - Q^{\star}(s_0, \hat{\pi}(s_0)) + Q^{\star}(s_0, \hat{\pi}(s_0)) - Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) \\ &\leq 2\varepsilon_Q + \mathbb{E}_{s_0} Q^{\star}(s_0, \hat{\pi}(s_0)) - Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) \end{split}$$

Here the main inequality uses the fact that  $\hat{\pi}$  is greedy with respect to  $Q^{(T)}$ , so  $Q^{(T)}(s, \hat{\pi}(s)) \geq Q^{(T)}(s, \pi^{\star}(s))$ . Now, we may unroll the last term here since  $Q^{\star}(s_0, \hat{\pi}(s_0)) = r(s_0, \hat{\pi}(s_0)) + \gamma \mathbb{E}_{s_1 \sim P(s_0, \hat{\pi}(s_0))} Q^{\star}(s_1, \pi^{\star}(s_1))$  while  $Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) = r(s_0, \hat{\pi}(s_0)) + \gamma \mathbb{E}_{s_1 \sim P(s_0, \hat{\pi}(s_0))} Q^{\hat{\pi}}(s_1, \hat{\pi}(s_1))$ . This gives

$$\mathbb{E}_{s_0} Q^{\star}(s_0, \hat{\pi}(s_0)) - Q^{\hat{\pi}}(s_0, \hat{\pi}(s_0)) = \gamma \mathbb{E}_{s_1 \sim d_1^{\hat{\pi}}} Q^{\star}(s_1, \pi^{\star}(s_1)) - Q^{\hat{\pi}}(s_1, \hat{\pi}(s_1)),$$

which has the same form as what we started with. Thus by unrolling, we obtain the bound:

$$J(\pi^{\star}) - J(\hat{\pi}) \le \frac{2\varepsilon_Q}{1 - \gamma}$$

4. **Bellman rank.** The key calculation is that in a linear MDP, the Bellman backup of any function  $g: \mathcal{S} \to \mathbb{R}$  is *linear* in the true features  $\phi^*$ . To see this, consider some policy  $\pi$  and note that

$$\mathbb{E}_{s_h \sim d_h^{\pi}} g(s_h) = \mathbb{E}_{s_{h-1}, a_{h-1} \sim d_{h-1}^{\pi}} \int P(s_h \mid s_{h-1}, a_{h-1}) g(s_h) ds_h$$

$$= \mathbb{E}_{s_{h-1}, a_{h-1} \sim d_{h-1}^{\pi}} \int \langle \phi^{\star}(s_{h-1}, a_{h-1}), \mu^{\star}(s_h) \rangle g(s_h) ds_h$$

$$= \left\langle \mathbb{E}_{s_{h-1}, a_{h-1} \sim d_{h-1}^{\pi}} \phi^{\star}(s_{h-1}, a_{h-1}), \int \mu^{\star}(s_h) g(s_h) ds_h \right\rangle$$

This immediately shows that the Bellman error  $\mathcal{E}_h(\pi, f)$  factorizes, since we can take g in the above derivation to be  $g: s_h \mapsto \mathbb{E}_{a_h \sim \pi_f(s_h)}[(f - \mathcal{T}f)(s_h, a_h)]$ , which is only a function of the state  $s_h$ . Then

$$\mathcal{E}_h(\pi, f) = \mathbb{E}_{s_h \sim d_h^{\pi}} \mathbb{E}_{a_h \sim \pi_f(s_h)} \left[ (f - \mathcal{T}f)(s_h, a_h) \right] = \mathbb{E}_{s_h \sim d_h^{\pi}} \left[ g(s_h) \right]$$
$$= \left\langle \mathbb{E}_{s_{h-1}, a_{h-1} \sim d_{h-1}^{\pi}} \phi^{\star}(s_{h-1}, a_{h-1}), \int \mu^{\star}(s_h) g(s_h) ds_h \right\rangle$$

Thus, we can take  $w_h(\pi) = \mathbb{E}_{s_{h-1}, a_{h-1} \sim d_{h-1}^{\pi}} \phi^{\star}(s_{h-1}, a_{h-1})$  and we can take  $v_h(f) = \int \mu^{\star}(s_h) \mathbb{E}_{a \sim \pi_f(s_h)}[(f - \mathcal{T}f)(s_h, a_h)] ds_h$  and see that the Bellman rank is d. Note that these embeddings also satisfy reasonable normalization conditions, since we typically assume  $\|\phi_h^{\star}\|_2$  is bounded, and it is natural to assume that both f and  $\mathcal{T}f$  are bounded as well.