Log-Linear Models a.k.a. Logistic Regression, Maximum Entropy Models

Introduction to Natural Language Processing Computer Science 585—Fall 2009 University of Massachusetts Amherst

David Smith (some slides from Jason Eisner and Dan Klein)

Probability is Useful

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 - Bayesian smoothing: max $p(\theta|data) = max p(\theta, data) = p(\theta)p(data|\theta)$

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- p(...) has to capture our intuitions about the ling. data



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 - Possible parses (or whatever) have scores.
 - Pick the one with the best score.
 - How do you define the score?
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Note: Today we'll use +logprob not –logprob: i.e., bigger weights are better.

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- PCFG: log p(NP VP | S) + log p(Papa | NP) + log p(VP PP | VP) ...
 - Can regard any linguistic object as a collection of features (here, tree = a collection of context-free rules)
 - Weight of the object = total weight of features
 - Our weights have always been conditional log-probs (≤ 0)
 - but that is going to change in a few minutes!
- HMM tagging: ... + log p(t7 | t5, t6) + log p(w7 | t7) + ...
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- Multiply indep. conditional probs normalized, unlike scores
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Some useful features:

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- Contains supercalifragilistic
- Contains a dollar amount under \$100
- Contains an imperative sentence
- Reading level = 8th grade
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spam ting
.5 .02
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 but here are the emails with both features – only 25x!/
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span ting
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log prob

Mentions money

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- Well, maybe we can add up scores and <u>pretend</u> like we got a log probability:

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- But not clear how to restructure these features like that: total: 5.77
 - Contains Buy

+4

+1

+2

-3

+5

- +0.2**Contains** supercalifragilistic
 - Contains a dollar amount under \$100
 - Contains an imperative sentence
 - Reading level = 7th grade
 - Mentions money (use word classes and/or regexp to detect this)
 - Boy, we'd like to be able to throw all that useful stuff in without worrying about feature overlap/independence.
 - Well, maybe we can add up scores and pretend like we got a log probability: log p(feats | spam) = 5.77

- Naïve Bayes needs overlapping but independent features
- But not clear how to restructure these features like that: total: 5.77
 - Contains Buy

+4

+1

+2

-3

+5

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Renormalize by 1/Z to get a

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 p(m | spam) = (1/Z(λ)) exp ∑_i λ_i f_i(m) where m is the email message
 λ_i is weight of feature i
 f_i(m)∈{0,1} according to whether m has feature i

More generally, allow $f_i(m) = count$ or strength of feature.

 $1/Z(\lambda)$ is a normalizing factor making $\sum_{m} p(m \mid spam)=1$ (summed over all possible messages m! hard to find!)



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- But we can fix this ...

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 - $\prod_j p(c_j \mid m_j)$ is still convex, so easy to maximize too

Generative vs. Conditional

- What is the most likely label for a given input?
- How likely is a given label for a given input?
- What is the most likely input value?
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- OUCH!

	A	В	С	D	E	F	G	Н	Ι	J
Buy	0.051	0.003	0.029	0.003	0.003	0.003	0.003	0.003	0.003	0.003
Other	0.499	0.045	0.045	0.045	0.045	0.045	0.045	0.045	0.045	0.045

Column A sums to 0.55 ("55% of all messages are in class A")

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- Row Buy sums to 0.1
- (Buy, A) and (Buy, C) cells sum to 0.08 ("80% of the 10%")

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 Entropy = -.051 log .051 - .0025 log .0025 - .029 log .029 - ...
 Largest if probabilities are evenly distributed

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- Given these constraints, fill in cells "as equally as possible": maximize the entropy
- Now p(Buy, C) = .029 and p(C | Buy) = .29
- We got a compromise: p(C | Buy) < p(A | Buy) < .55</p>

Generalizing to More Features



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 - Gives another motivation for our log-linear approach.

Log-linear form derivation

Say we are given some *constraints* in the form of feature expectations:

$$\sum_x p(x) f_i(x) = \alpha_i$$

- In general, there may be many distributions p(x) that satisfy the constraints. Which one to pick?
- The one with maximum entropy (making fewest possible additional assumptions---Occum's Razor)
- This yields an optimization problem

$$\label{eq:max} \begin{split} \max H(p(x)) &= -\sum_x p(x)\log p(x)\\ \text{Subject to } \sum_x p(x)f_i(x) = \alpha_i, \forall i \ \text{and} \ \sum_x p(x) = 1 \end{split}$$

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Log-linear form derivation

• To solve the maxent problem, we use Lagrange multipliers:

$$\begin{split} L &= -\sum_{\mathbf{x}} p(\mathbf{x}) \log p(\mathbf{x}) - \sum_{i} \theta_{i} \left(\sum_{\mathbf{x}} p(\mathbf{x}) f_{i}(\mathbf{x}) - \alpha_{i} \right) - \mu \left(\sum_{\mathbf{x}} p(\mathbf{x}) - 1 \right) \\ \frac{\partial L}{\partial p(\mathbf{x})} &= 1 + \log p(\mathbf{x}) - \sum_{i} \theta_{i} f_{i}(\mathbf{x}) - \mu \\ p^{*}(\mathbf{x}) &= e^{\mu - 1} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\} \\ Z(\theta) &= e^{1 - \mu} = \sum_{\mathbf{x}} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\} \\ p(\mathbf{x}|\theta) &= \frac{1}{Z(\theta)} \exp \left\{ \sum_{i} \theta_{i} f_{i}(\mathbf{x}) \right\} \end{split}$$

- So feature constraints + maxent implies exponential family.
- Problem is convex, so solution is unique.

MaxEnt = Max Likelihood

Define two submanifolds on the probability simplex $p(\mathbf{x})$.

The first is \mathcal{E} , the set of all exponential family distributions based on a particular set of features $f_i(\mathbf{x})$.

The second is \mathcal{M} , the set of all distributions that satisfy the feature expectation constraints.

They intersect at a single distribution p_M , the maxent, maximum likelihood





- Maximum Likelihood (Conditional) Models :
 - Given a model form, choose values of parameters to maximize the (conditional) likelihood of the data.
- Exponential model form, for a data set (C,D):

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} P(c \mid d, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}$$


Building a Maxent Model

- Define features (indicator functions) over data points.
 - Features represent sets of data points which are distinctive enough to deserve model parameters.
 - Usually features are added incrementally to "target" errors.
- For any given feature weights, we want to be able to calculate:
 - Data (conditional) likelihood
 - Derivative of the likelihood wrt each feature weight
 - Use expectations of each feature according to the model
- Find the optimum feature weights (next part).

 The (log) conditional likelihood is a function of the iid data (C,D) and the parameters λ:

$$\log P(C \mid D, \lambda) = \log \prod_{(c,d) \in (C,D)} P(c \mid d, \lambda) = \sum_{(c,d) \in (C,D)} \log P(c \mid d, \lambda)$$

If there aren't many values of c, it's easy to calculate:

The Likelihood Value

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c,d)}$$

We can separate this into two components:

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \exp \sum_{i} \lambda_i f_i(c,d) - \sum_{(c,d) \in (C,D)} \log \sum_{c'} \exp \sum_{i} \lambda_i f_i(c',d)$$

$$\log P(C \mid D, \lambda) = N(\lambda) - M(\lambda)$$

The derivative is the difference between the derivatives of each component



$$\frac{\partial N(\lambda)}{\partial \lambda_{i}} = \frac{\partial \sum_{(c,d)\in(C,D)} \log \exp \sum_{i} \lambda_{ci} f_{i}(c,d)}{\partial \lambda_{i}} = \frac{\partial \sum_{(c,d)\in(C,D)} \sum_{i} \lambda_{i} f_{i}(c,d)}{\partial \lambda_{i}}$$
$$= \sum_{(c,d)\in(C,D)} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c,d)}{\partial \lambda_{i}}$$
$$= \sum_{(c,d)\in(C,D)} f_{i}(c,d)$$

Derivative of the numerator is: the empirical count(f_i , c)

The Derivative II: Denominator

$$\begin{aligned} \frac{\partial M(\lambda)}{\partial \lambda_{i}} &= \frac{\partial \sum_{(c,d)\in(C,D)} \log \sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}} \\ &= \sum_{(c,d)\in(C,D)} \frac{1}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \frac{\partial \sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}} \\ &= \sum_{(c,d)\in(C,D)} \frac{1}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \sum_{c'} \frac{\exp \sum_{i} \lambda_{i} f_{i}(c',d)}{1} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}} \\ &= \sum_{(c,d)\in(C,D)} \sum_{c'} \frac{\exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}} \\ &= \sum_{(c,d)\in(C,D)} \sum_{c'} \frac{\exp \sum_{i} \lambda_{i} f_{i}(c',d)}{\sum_{c''} \exp \sum_{i} \lambda_{i} f_{i}(c'',d)} \frac{\partial \sum_{i} \lambda_{i} f_{i}(c',d)}{\partial \lambda_{i}} \\ &= \sum_{(c,d)\in(C,D)} \sum_{c'} P(c'|d,\lambda) f_{i}(c',d) = \text{predicted count}(f_{i},d) \end{aligned}$$

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The Derivative III

 $\frac{\partial \log P(C \mid D, \lambda)}{\partial \lambda_i} = \operatorname{actual count}(f_i, C) - \operatorname{predicted count}(f_i, \lambda)$

- The optimum parameters are the ones for which each feature's predicted expectation equals its empirical expectation. The optimum distribution is:
 - Always unique (but parameters may not be unique)
 - Always exists (if features counts are from actual data).
- Features can have high model expectations (predicted counts) either because they have large weights or because they occur with other features which have large weights.



We have a function to optimize:

$$\log P(C \mid D, \lambda) = \sum_{(c,d) \in (C,D)} \log \frac{\exp \sum_{i} \lambda_{i} f_{i}(c,d)}{\sum_{c'} \exp \sum_{i} \lambda_{i} f_{i}(c,d)}$$

- We know the function's derivatives: $\partial \log P(C \mid D, \lambda) / \partial \lambda_i = \operatorname{actual count}(f_i, C) - \operatorname{predicted count}(f_i, \lambda)$
- Perfect situation for general optimization (Part II)
 By gradient ascent or conjugate gradient.



Comparison to Naïve-Bayes

 The primary differences between Naïve-Bayes and maxent models are:

Naïve-Bayes

Trained to maximize joint likelihood of data and classes. Features assumed to supply independent evidence. Feature weights can be set independently.

Features must be of the conjunctive $\Phi(d) \wedge c = c_i$ form.

Maxent

Trained to maximize the conditional likelihood of classes. Features weights take feature dependence into account. Feature weights must be mutually estimated.

Features need not be of the conjunctive form (but usually are).

Overfitting

If we have too many features, we can choose weights to model the training data perfectly.

- If we have a feature that only appears in spam training, not ling training, it will get weight ∞ to maximize p(spam | feature) at 1.
- These behaviors overfit the training data.
- Will probably do poorly on test data.

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 - Require every feature to occur > 4 times, and > 0 times with ling, and > 0 times with spam.

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- 3. Smooth the observed feature counts.
- 4. Smooth the weights by using a prior.
 - max $p(\lambda | data) = max p(\lambda, data) = p(\lambda)p(data | \lambda)$
 - decree $p(\lambda)$ to be high when most weights close to 0



Smoothing: Priors (MAP)

- What if we had a prior expectation that parameter values wouldn't be very large?
- We could then balance evidence suggesting large parameters (or infinite) against our prior.
- The evidence would never totally defeat the prior, and parameters would be smoothed (and kept finite!).
- We can do this explicitly by changing the optimization objective to maximum posterior likelihood:

 $\log P(C, \lambda \mid D) = \log P(\lambda) + \log P(C \mid D, \lambda)$ Posterior Prior Evidence



- Gaussian, or quadratic, priors:
 - Intuition: parameters shouldn't be large.
 - Formalization: prior expectation that each parameter will be distributed according to a gaussian with mean μ and variance σ².

$$P(\lambda_i) = \frac{1}{\sigma_i \sqrt{2\pi}} \exp\left(-\frac{(\lambda_i - \mu_i)^2}{2\sigma_i^2}\right)$$

- Penalizes parameters for drifting to far from their mean prior value (usually µ=0).
- $2\sigma^2 = 1$ works surprisingly well.



Recipe for a Conditional MaxEnt Classifier

1. Gather *constraints* from training data:

$$\alpha_{iy} = \tilde{E}[f_{iy}] = \sum_{x_j, y_j \in D} f_{iy}(x_j, y_j)$$

- 2. Initialize all parameters to zero.
- 3. Classify training data with current parameters. Calculate *expectations*. $E_{\Theta}[f_{iy}] = \sum_{x_j \in D} \sum_{y'} p_{\Theta}(y'|x_j) f_{iy}(x_j, y')$
- 4. Gradient is $\tilde{E}[f_{iy}] E_{\Theta}[f_{iy}]$
- 5. Take a step in the direction of the gradient
- 6. Until convergence, return to step 3.